

## INADMISSIBILITY OF THE SAMPLE VARIANCE

For any observations  $X_1, \dots, X_n$  with  $n \geq 2$ , assumed to be i.i.d. with finite variance  $\sigma^2$ , let  $s_X^2 = \frac{1}{n-1} \sum_{j=1}^n (X_j - \bar{X})^2$ , the usual unbiased estimator of  $\sigma^2$ . Y. G. Yatracos (2005) proved that  $s_X^2$  is inadmissible as an estimator of  $\sigma^2$  by providing an estimator which has smaller mean-squared error in all relevant cases. More precisely, he proved the following:

**Theorem 1** (Yatracos). *There is a constant  $c_n$  depending on  $n$ , namely  $c_n = \frac{(n+2)(n-1)}{n(n+1)}$ , such that for any  $n \geq 2$  and for all  $X_1, \dots, X_n$  i.i.d. with  $E(X_1^4) < +\infty$  and variance  $\sigma^2$  with  $\sigma > 0$ , the mean-square error of  $c_n s_X^2$  as an estimator of  $\sigma^2$  is less than that of  $s_X^2$ .*

*Proof.* First let's find the mean-square error as a function of an unspecified  $c_n$ . We can assume without loss of generality that  $\mu := EX_1 = 0$  because replacing all  $X_j$  by  $X_j - \mu$  does not change either  $\sigma^2$  or  $s_X^2$ . Then let  $\tau := \tau_X := E(X_1^4)/\sigma^4$ . We have

$$(1) \quad E((c_n s_X^2 - \sigma^2)^2) = c_n^2 E(s_X^4) - 2c_n \sigma^4 + \sigma^4,$$

and

$$(2) \quad E(s_X^4) := E\left(\left(s_X^2\right)^2\right) = \frac{1}{(n-1)^2} [n(n-1)T_{12} + nT_4]$$

where

$$T_{12} := E\{(X_1 - \bar{X})^2(X_2 - \bar{X})^2\}$$

and

$$T_4 := E\{(X_1 - \bar{X})^4\}.$$

Next,

$$\begin{aligned} T_{12} &= E[(X_1^2 - 2X_1\bar{X} + \bar{X}^2)(X_2^2 - 2X_2\bar{X} + \bar{X}^2)] \\ &= \sigma^4 - 2E(X_1X_2^2\bar{X}) - 2E(X_2X_1^2\bar{X}) + 4E(X_1X_2\bar{X}^2) \\ &\quad - 2E(X_1\bar{X}^3) - 2E(X_2\bar{X}^3) + E(X_1^2\bar{X}^2) + E(X_2^2\bar{X}^2) + E(\bar{X})^4 \\ &= \sigma^4 - 4E(X_1X_2^2\bar{X}) + 4E(X_1X_2\bar{X}^2) - 4E(X_1\bar{X}^3) + 2E(X_1^2\bar{X}^2) + E(\bar{X})^4 \\ &= \sigma^4 - 4\frac{\sigma^4}{n} + 8\frac{\sigma^4}{n^2} - 4\left[\frac{\tau\sigma^4 + 3(n-1)\sigma^4}{n^3}\right] \\ &\quad + 2\left[\frac{\tau\sigma^4 + (n-1)\sigma^4}{n^2}\right] + \frac{n\tau\sigma^4 + 3n(n-1)\sigma^4}{n^4}, \end{aligned}$$

where  $3n(n-1) = \binom{n}{2} \binom{4}{2}$ . Thus

$$\begin{aligned}
 T_{12} &= \sigma^4 \left\{ \left[ 1 - \frac{4}{n} + \frac{8}{n^2} - 12 \frac{(n-1)}{n^3} + \frac{2(n-1)}{n^2} + \frac{3(n-1)}{n^3} \right] \right. \\
 &\quad \left. + \tau \left[ -\frac{4}{n^3} + \frac{2}{n^2} + \frac{1}{n^3} \right] \right\} \\
 (3) \quad &= \sigma^4 \left\{ 1 - \frac{2}{n} - \frac{3}{n^2} + \frac{9}{n^3} + \tau \left[ \frac{2}{n^2} - \frac{3}{n^3} \right] \right\}.
 \end{aligned}$$

That finishes the representation of  $T_{12}$ . For another term of (2) we have

$$\begin{aligned}
 T_4 &= E(X_1^4) - 4E(X_1^3 \bar{X}) + 6E(X_1^2 \bar{X}^2) - 4E(X_1 \bar{X}^3) + E(\bar{X}^4) \\
 &= \sigma^4 \left( \tau - 4 \frac{\tau}{n} + 6 \left\{ \frac{\tau + n - 1}{n^2} \right\} - 4 \left\{ \frac{\tau + 3(n-1)}{n^3} \right\} + \frac{\tau + 3(n-1)}{n^3} \right) \\
 &= \sigma^4 \left[ \frac{6(n-1)}{n^2} - \frac{9(n-1)}{n^3} + \tau \left\{ 1 - \frac{4}{n} + \frac{6}{n^2} - \frac{4}{n^3} + \frac{1}{n^3} \right\} \right] \\
 (4) \quad &= \sigma^4 \left[ \tau \left\{ 1 - \frac{4}{n} + \frac{6}{n^2} - \frac{3}{n^3} \right\} + \frac{6}{n} - \frac{15}{n^2} + \frac{9}{n^3} \right].
 \end{aligned}$$

Inserting the results of (3) and (4) into (2) gives

$$E[(c_n s_X^2 - \sigma^2)^2] = \sigma^4 [c_n^2 A_n - 2c_n + 1]$$

where

$$\begin{aligned}
 A_n &= \frac{n}{n-1} \left[ 1 - \frac{2}{n} - \frac{3}{n^2} + \frac{9}{n^3} + \tau \left( \frac{2}{n^2} - \frac{3}{n^3} \right) \right] \\
 &\quad + \frac{n}{(n-1)^2} \left[ \frac{6}{n} - \frac{15}{n^2} + \frac{9}{n^3} + \tau \left( 1 - \frac{4}{n} + \frac{6}{n^2} - \frac{3}{n^3} \right) \right].
 \end{aligned}$$

This implies

$$(5) \quad E[(c_n s_X^2 - \sigma^2)^2] = \sigma^4 \left[ \frac{c_n^2}{(n-1)^2} (B_n + \tau C_n) - 2c_n + 1 \right]$$

where

$$B_n = n(n-1) - 2(n-1) - \frac{3(n-1)}{n} + \frac{9(n-1)}{n^2} + 6 - \frac{15}{n} + \frac{9}{n^2}$$

and

$$C_n = \frac{2(n-1)}{n} - \frac{3(n-1)}{n^2} + n - 4 + \frac{6}{n} - \frac{3}{n^2}.$$

Then

$$B_n = n^2 - 3n + 2 - 3 + 6 + \frac{1}{n}(3 + 9 - 15) + \frac{1}{n^2}(-9 + 9)$$

$$= n^2 - 3n + 5 - \frac{3}{n}$$

and

$$C_n = n - 4 + 2 + \frac{1}{n}(6 - 2 - 3) + \frac{1}{n^2}(3 + 11) = n - 2 + \frac{1}{n} + \frac{14}{n^2}.$$

Let

$$D_n := n^2 - 3n + 5 - \frac{3}{n} + \tau \left( n - 2 + \frac{1}{n} + \frac{14}{n^2} \right).$$

Thus the mean-square error (MSE) on the left in (1) or (5) equals

$$\sigma^4 \left[ \frac{c_n^2}{(n-1)^2} D_n - 2c_n + 1 \right].$$

So, in order to get a smaller MSE for a given  $c_n < 1$  than for  $c_n = 1$ , we ask whether

$$\frac{c_n^2}{(n-1)^2} D_n - 2c_n + 1 < \frac{D_n}{(n-1)^2} - 1,$$

the right side being the result if  $c_n$  on the left is replaced by 1. in other words whether

$$\frac{c_n^2 - 1}{(n-1)^2} D_n < 2c_n - 2,$$

or, reversing signs and dividing by  $1 - c_n$ , whether

$$\frac{1 + c_n}{(n-1)^2} D_n > 2.$$

We can also write

$$D_n = n^2 - 2n + 3 - \frac{2}{n} + \frac{14}{n^2} + (\tau - 1) \left( n - 2 + \frac{1}{n} + \frac{14}{n^2} \right).$$

Now  $n - 2 + \frac{1}{n} + \frac{14}{n^2} > 0$  for all  $n \geq 2$  because  $n^3 - 2n^2 + n + 14 = n^2(n-2) + n + 14 > 0$ . Thus

$$D_n \geq E_n := n^2 - 2n + 3 - \frac{2}{n} + \frac{14}{n^2}$$

since  $\tau \geq 1$ . Now plugging in Yatracos's value  $c_n = \frac{(n+2)(n-1)}{n(n+1)}$  we're asking whether

$$\left[ 1 + \frac{(n+2)(n-1)}{n(n+1)} \right] E_n > 2(n-1)^2,$$

or equivalently whether

$$[n(n+1) + (n+2)(n-1)] E_n > 2(n-1)^2 n(n+1),$$

or whether

$$(2n^2 + 2n - 2) E_n > 2(n^4 - n^3 - n^2 + n),$$

or whether

$$n^4 - n^3 + n(-2 + 3 + 2) + 14 - 2 - 3 + \frac{1}{n}(14 + 2) - \frac{14}{n^2} > n^4 - n^3 - n^2 + n.$$

This reduces to

$$n^2 + 2n + 9 + \frac{16}{n} - \frac{14}{n^2} > 0$$

which is clearly true for all all  $n \geq 2$ , proving the theorem.  $\square$

**Remarks.** It do not know at this writing whether with Yatracos's  $c_n$  the resulting estimator is admissible.

The mean-square error of  $c_n s_X^2$  depends on the distribution of  $X_1$  only through the number  $\tau$ . For each  $\tau \geq 1$ , it seems possible to determine a  $c_n = c_n(\tau)$  which minimizes the mean-square error.

**Notes.** I appreciate that Y. Yatracos sent me an advance copy, in 2004 or 2005, of his work. In 18.466, Fall 2005, Songzi Du wrote a term paper in which he gave an exposition of Yatracos's paper, with proofs, which helped my understanding. Both Yatracos's paper, and Songzi Du's exposition, use a method of augmented samples. Whereas, I looked for a direct proof of Yatracos's theorem without augmented samples.

#### REFERENCES

Du, Songzi (2005). Mean-squared error reduction for unbiased estimators. Term paper, 18.466, Fall 2005.

Yatracos, Yannis G. (2005). Artificially augmented samples, and mean squared error reduction. *J. Amer. Statistic. Assoc.* **100**, pp. 1168–1175.