

Collision time for Markov chains

Lemma 0.1. *For a given reversible n -state Markov chain P and its stationary distribution π ($\pi_i = O(1/n), \forall i$), denote X_t ($t < \sqrt{n}$) be the t^{th} state of the random walk according to P . Consider a fixed starting state $X_0 = s$, then*

$$\Pr\left[\bigcap_{0 < i < t} X_i \neq s\right] \leq 1 - \Theta(t\pi_s).$$

Proof. At first, claim

$$\Pr[X_{2t} = s] \geq \pi_s$$

. It is because

$$\begin{aligned} \Pr[X_{2t} = s] &= \sum_i \Pr[X_t = i] \Pr[X_{2t} = s | X_t = i] \\ &= \sum_i \Pr[X_t = i] \Pr[X_t = s | X_0 = i] \\ &= \sum_i P_{si}^t P_{is}^t \\ &= \sum_i P_{si}^t P_{si}^t \frac{\pi_s}{\pi_i} \\ &= \pi_s \sum_i (P_{si}^t)^2 \frac{1}{\pi_i} \\ &= \pi_s \left(\sum_i (P_{si}^t)^2 \frac{1}{\pi_i} \right) \left(\sum_i \pi_i \right) \\ &\geq \pi_s \left(\sum_i P_{si}^t \right)^2 \\ &= \pi_s \end{aligned}$$

$\Pr[\bigcap_{0 < i < t} X_i \neq s] = 1 - \Pr[\bigcup_{0 < i < t} X_i = X_0]$, and we divide two cases.

Case 1. $1/2 < \sum_{0 < i < t} \Pr[X_0 = X_i]$

It means $\exists j, \Pr[X_0 = X_j] > \frac{1}{2t}$. Hence,

$$1 - \Pr\left[\bigcup_{0 < i < t} X_0 = X_i\right] < 1 - \Pr[X_0 = X_j] < 1 - \frac{1}{2t} < 1 - \frac{t}{2n} < 1 - \Theta(t\pi_s).$$

The third inequality is because $t < \sqrt{n}$.

Case 2. $1/2 > \sum_{0 < i < t} \Pr[X_0 = X_i] = T$

$$\begin{aligned}
& 1 - \Pr\left[\bigcup_{0 < i < t} X_0 = X_i\right] \\
& < 1 - \sum_{0 < i < t} \Pr[X_0 = X_i] + \sum_{0 < i < j < t} \Pr[X_0 = X_i \& X_0 = X_j] \\
& < 1 - \sum_{0 < i < t} \Pr[X_0 = X_i] + \sum_{0 < i < j < t} \Pr[X_0 = X_i \& X_i = X_j] \\
& < 1 - \sum_{0 < i < t} \Pr[X_0 = X_i] + \sum_{0 < i < j < t} \Pr[X_0 = X_i] \Pr[X_0 = X_{j-i}] \\
& < 1 - T + T^2 \\
& < 1 - T/2
\end{aligned}$$

The last inequality is because $T < 1/2$. By the above claim, we know $T > \Theta(t\pi_s)$, which implies $1 - T/2 < 1 - \Theta(t\pi_s)$ and complete the proof. \square

Theorem 0.2. *The collision time of the random walk induced by n -state reversible Markov chain P that has the uniform stationary distribution is $O(\sqrt{n})$.*

Proof. Let X_t be the state after t steps. Define the following event A_i, B_i, C_i :

$$\begin{aligned}
A_i &= \bigcap_{0 \leq i < j \leq \sqrt{n}} X_i \neq X_j \\
B_i &= \bigcap_{0 \leq j < k \leq i} X_j \neq X_k \\
C_i &= \bigcap_{0 \leq j \leq i < k \leq \sqrt{n}} X_j \neq X_k
\end{aligned}$$

Our goal is to show the following probability Q is bounded by a constant < 1 :

$$\begin{aligned}
Q &= \Pr\left[\bigcap_{0 \leq i < j \leq \sqrt{n}} X_i \neq X_j\right] \\
&= \prod_{i=0}^{\sqrt{n}} \Pr\left[\bigcap_{j \leq i} X_j \neq X_{i+1} \mid B_i\right] \\
&= \prod_{i=0}^{\sqrt{n}} E\left[\sum_{k \neq X(t), \forall t \leq i} P_{X_i k} \mid B_i\right]
\end{aligned}$$

Now suppose $\forall i, E\left[\sum_{k \neq X(t), \forall t \leq i} P_{X_i k} \mid B_i\right] > 1/2$, otherwise there is nothing to prove.

$$\begin{aligned}
Q &= \Pr\left[\bigcap_{1 \leq i < \sqrt{n}} A_i\right] \\
&= \prod_{1 \leq i < \sqrt{n}} \Pr[A_i \mid A_1, \dots, A_{i-1}]
\end{aligned}$$

Also, $Pr[A_i|A_1, \dots, A_{i-1}]$ is same as $Pr[A_i]$ in the Markov chain $P^{(i)} = P - \{X_0, \dots, X_{i-1}\}$, which is also reversible. Let $\pi^{(i)}$ be the stationary distribution of $P^{(i)}$. By Lemma 0.1,

$$Pr[A_i|A_1, \dots, A_{i-1}] \leq E \left[1 - \Theta \left((\sqrt{n} - i) \pi_{X_i}^{(i)} \right) \mid A_1, \dots, A_{i-1} \right]$$

Note that

$$\pi_{X_i}^{(i)} \approx \frac{\sum_{k \neq X(t), \forall t \leq i} P_{X_i k}}{n}.$$

Because $\bigcap_{j < i} A_j = B_i \cap C_i$, $E[\pi_{X_i}^{(i)}|A_1, \dots, A_{i-1}] = E[\pi_{X_i}^{(i)}|B_i, C_i]$.

$$\begin{aligned} E[\pi_{X_i}^{(i)}|B_i] &= \Pr(C_i|B_i)E[\pi_{X_i}^{(i)}|B_i, C_i] + \Pr(\bar{C}_i|B_i)E[\pi_{X_i}^{(i)}|B_i, \bar{C}_i] \\ &\leq \Pr(C_i|B_i)E[\pi_{X_i}^{(i)}|B_i, C_i] + \Pr(\bar{C}_i|B_i)\frac{1}{n} \\ &\leq E[\pi_{X_i}^{(i)}|B_i, C_i] + \frac{1}{100n} \end{aligned}$$

The last inequality is because $\Pr(C_i|B_i) \geq 99/100$, otherwise there is nothing to prove because $Q < \Pr(B_i, C_i) = \Pr(B_i) \Pr(C_i|B_i) < 99/100$.

$$\begin{aligned} \sum_{i \leq \sqrt{n}} (\sqrt{n} - i) E[\pi_{X_i}^{(i)}|B_i, C_i] &\geq \sum_{i \leq \sqrt{n}} (\sqrt{n} - i) \left(E[\pi_{X_i}^{(i)}|B_i] - \frac{1}{100n} \right) \\ &\geq \sum_{i \leq \sqrt{n}} (\sqrt{n} - i) \left(\frac{1}{2n} - \frac{1}{100n} \right) \\ &\geq 0.1 \end{aligned}$$

Therefore,

$$\begin{aligned} Q &= \prod_{1 \leq i < \sqrt{n}} Pr[A_i|A_1, \dots, A_{i-1}] \\ &\leq \prod_{1 \leq i < \sqrt{n}} E \left[1 - \Theta \left((\sqrt{n} - i) \pi_{X_i}^{(i)} \right) \mid A_1, \dots, A_{i-1} \right] \\ &\approx \text{Exp} \left[- \sum_{i \leq \sqrt{n}} (\sqrt{n} - i) E[\pi_{X_i}^{(i)}|B_i, C_i] \right] \\ &\leq e^{-0.1} \end{aligned}$$

□