

Matrix Models for Beta Ensembles

Ioana Dumitriu* and Alan Edelman†

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Abstract

This paper constructs tridiagonal random matrix models for general ($\beta > 0$) β -Hermite (Gaussian) and β -Laguerre (Wishart) ensembles. These generalize the well-known Gaussian and Wishart models for $\beta = 1, 2, 4$. Furthermore, in the cases of the β -Laguerre ensembles, we eliminate the exponent quantization present in the previously known models.

We further discuss applications for the new matrix models, and present some open problems.

*Department of Mathematics, room 2-331, Massachusetts Institute of Technology, 77 Massachusetts Avenue, Cambridge, MA 02139 (dumitriu@math.mit.edu)

†Department of Mathematics, room 2-388, Massachusetts Institute of Technology, 77 Massachusetts Avenue, Cambridge, MA 02139 (edelman@math.mit.edu; <http://math.mit.edu/~edelman>)

1 Introduction

1.1 Overview

Classical Random Matrix Theory focuses on the random matrix models in the following 3×3 table:

	Real, $\beta = 1$	Complex, $\beta = 2$	Quaternion, $\beta = 4$
Hermite	GOE	GUE	GSE
Laguerre	Real Wishart	Complex Wishart	(Quaternion Wishart)
Jacobi	Real MANOVA	Complex MANOVA	(Quaternion MANOVA)

The two entries in parentheses (in the third column) correspond to less-studied random matrix models; the others are mainstream and have been extensively researched and publicized. The three columns correspond to Dyson’s “threefold way” $\beta = 1, 2,$ and 4 ; the three rows correspond to the weight function associated to the random matrix model. Other weight functions have also been considered (for example, the uniform weight on the unit circle corresponds to the circular ensembles).

Zirnbauer [33] and Ivanov [12] produced a more general taxonomy of random matrix models. Their characterizations (“tenfold”, respectively “twelvefold”) are based on symmetric spaces, and include Hermite, Laguerre, and Jacobi cases, and also the circular ensembles (each of their models can be associated with $\beta = 1, 2$ or 4).

We propose a random matrix program of study that would generalize β beyond the above threefold way, thus generalizing the 3×3 cartesian product to $3 \times \infty$, making the leap from discrete characterizations to continuous ones. A step in this direction has been initiated by Forrester [2], [10], who studied the β -ensembles in connection with multivariate orthogonal polynomials and Calogero-Sutherland-type quantum systems. Furthermore, in the case of the classical Laguerre and Jacobi models, our program goes beyond the quantized exponents forced by the classical models, and proposes continuous ones.

For the benefit of the reader we have expanded the 3×3 table with detailed information in Figure 1.

1.2 Background

The Gaussian (or Hermite) ensembles arise in physics, and are identified by Dyson [7] by the group over which they are invariant: Gaussian Orthogonal or for short GOE (with real entries), Gaussian Unitary or GUE (with complex entries), and Gaussian Symplectic or GSE (with quaternion entries). The Wishart ensembles arise in statistics, and the three corresponding models could be named Wishart real, Wishart complex, and Wishart quaternion.

The three Gaussian Ensembles have joint eigenvalue probability density function

$$\text{HERMITE:} \quad f_{\beta}(\lambda) = c_H^{\beta} \prod_{i < j} |\lambda_i - \lambda_j|^{\beta} e^{-\sum_{i=1}^n \lambda_i^2/2}, \quad (1)$$

with $\beta = 1$ corresponding to the reals, $\beta = 2$ to the complexes, $\beta = 4$ to the quaternions, and

Random Matrix Ensembles

Joint Eigenvalue Distributions: $c | \prod_{i=1}^n \prod_{j=1}^n V(\lambda_{ij})$

Technical name	Traditional name	Field	Property	Invariance	Connected Matrix Problem
Hermite ensembles $V(\square) = e^{-\square^2/2}$	Gaussian ensembles GOE GUE GSE	\mathbb{R} \mathbb{C} \mathbb{H}	Symmetric Hermitian Self-Dual	$A \leftrightarrow Q^T A Q$ $A \leftrightarrow U^H A U$ $A \leftrightarrow S^D A S$	(EIG) Symmetric Eigenvalue Problem
Laguerre ensembles $V(\square) = \square^a e^{-\square/2}$ $a = \frac{\square}{2}(n - m + 1) - 1$	Wishart ensembles Wishart real Wishart complex Wishart quaternion	\mathbb{R} \mathbb{C} \mathbb{H}	Positive Semi-Definite	$A \leftrightarrow Q^T A Q$ $A \leftrightarrow U^H A U$ $A \leftrightarrow S^D A S$	(SVD) Singular Value Decomposition
Jacobi ensembles $V(\square) = \prod_{i=1}^a (1 - \square_i)^b$ $a = \frac{\square}{2}(n_1 - m + 1) - 1$ $b = \frac{\square}{2}(n_2 - m + 1) - 1$	MANOVA ensembles MANOVA real MANOVA complex MANOVA quaternion	\mathbb{R} \mathbb{C} \mathbb{H}	Positive Semi-Definite	$X \leftrightarrow Q^T X Q_1$ $Y \leftrightarrow Q^T Y Q_1$ $X \leftrightarrow U^H X U_1$ $Y \leftrightarrow U^H Y U_1$ $X \leftrightarrow S^D X S_1$ $Y \leftrightarrow S^D Y S_1$	(QZ) Generalized Symmetric Eigenvalue Problem
Type of ensemble	MATLAB code				
Hermite	\square $A = \text{randn}(n)$; $A = (A+A')/2$; \square $A = \text{randn}(n) + i * \text{randn}(n)$; $A = (A+A')/2$; \square $X = \text{randn}(n) + i * \text{randn}(n)$; $Y = \text{randn}(n) + i * \text{randn}(n)$; \square $A = [X \ Y; -\text{conj}(Y) \ \text{conj}(X)]$; $A = (A+A')/2$; 				
Laguerre	\square $A = \text{randn}(m, n)$; $A = A * A'$; \square $A = \text{randn}(m, n) + i * \text{randn}(m, n)$; $A = A * A'$; \square $X = \text{randn}(m, n) + i * \text{randn}(m, n)$; $Y = \text{randn}(m, n) + i * \text{randn}(m, n)$; \square $A = [X \ Y; -\text{conj}(Y) \ \text{conj}(X)]$; $A = A * A'$; 				
Jacobi	\square $X = \text{randn}(m, n_1)$; $Y = \text{randn}(m, n_2)$; $A = (X * X') / (X * X' + Y * Y')$; \square $X = \text{randn}(m, n_1) + i * \text{randn}(m, n_1)$; $Y = \text{randn}(m, n_2) + i * \text{randn}(m, n_2)$; $A = (X * X') / (X * X' + Y * Y')$; \square $X_1 = \text{randn}(m, n_1) + i * \text{randn}(m, n_1)$; $X_2 = \text{randn}(m, n_1) + i * \text{randn}(m, n_1)$; \square $Y_1 = \text{randn}(m, n_2) + i * \text{randn}(m, n_2)$; $Y_2 = \text{randn}(m, n_2) + i * \text{randn}(m, n_2)$; \square $X = [X_1 \ X_2; -\text{conj}(X_2) \ \text{conj}(X_1)]$; $Y = [Y_1 \ Y_2; -\text{conj}(Y_2) \ \text{conj}(Y_1)]$; \square $A = (X * X') / (X * X' + Y * Y')$; 				

Figure 1: Random Matrix Ensembles. As a guide to MATLAB notation, $\text{randn}(m, n)$ produces an $m \times n$ matrix with i.i.d. standard normal entries, $\text{conj}(X)$ produces the complex conjugate of the matrix X , and the $'$ operator produces the conjugate transpose of a matrix. Also $[X \ Y; Z \ W]$ produces a 2×2 block matrix.

with

$$c_H^\beta = (2\pi)^{-n/2} \prod_{j=1}^n \frac{\Gamma(1 + \frac{\beta}{2})}{\Gamma(1 + \frac{\beta}{2}j)} . \quad (2)$$

The best references are Mehta [18] and the original paper by Dyson [7].

Similarly, the Wishart (or Laguerre) models have joint eigenvalue p.d.f.

$$\text{LAGUERRE:} \quad f_\beta(\lambda) = c_L^{\beta,a} \prod_{i<j} |\lambda_i - \lambda_j|^\beta \prod_i \lambda_i^{a-p} e^{-\sum_{i=1}^n \lambda_i/2} , \quad (3)$$

with $a = \frac{\beta}{2}n$ and $p = 1 + \frac{\beta}{2}(m-1)$. Again, $\beta = 1$ for the reals, $\beta = 2$ for the complexes, and $\beta = 4$ for the quaternions. The constant

$$c_L^{\beta,a} = 2^{-ma} \prod_{j=1}^m \frac{\Gamma(1 + \frac{\beta}{2})}{\Gamma(1 + \frac{\beta}{2}j)\Gamma(a - \frac{\beta}{2}(m-j))} . \quad (4)$$

Good references are [21], [8], and [13], and for $\beta = 4$, [17].

To complete the triad of classical orthogonal polynomials, we will mention the β -MANOVA ensembles, which are associated to the multivariate analysis of variance (MANOVA) model. They are better known in the literature as the Jacobi ensembles, with joint eigenvalue p.d.f.

$$\text{JACOBI:} \quad f_\beta(\lambda) = c_J^{\beta,a_1,a_2} \prod_{i<j} |\lambda_i - \lambda_j|^\beta \prod_{j=1}^n \lambda_i^{a_1-p} (1 - \lambda_i)^{a_2-p} , \quad (5)$$

with $a_1 = \frac{\beta}{2}n_1$, $a_2 = \frac{\beta}{2}n_2$, and $p = 1 + \frac{\beta}{2}(m-1)$. As usual, $\beta = 1$ for real and $\beta = 2$ for complex; also

$$c_J^{\beta,a_1,a_2} = \prod_{j=1}^m \frac{\Gamma(1 + \frac{\beta}{2})\Gamma(a_1 + a_2 - \frac{\beta}{2}(m-j))}{\Gamma(1 + \frac{\beta}{2}j)\Gamma(a_1 - \frac{\beta}{2}(m-j))\Gamma(a_2 - \frac{\beta}{2}(m-j))} . \quad (6)$$

The MANOVA real and complex cases ($\beta = 1$ and 2) have been studied by statisticians (see [21]).

Though ‘‘Gaussian’’, ‘‘Wishart’’, and ‘‘MANOVA’’ are the traditional names for the three types of β -ensembles, we prefer the sometimes used and technically more informative names ‘‘Hermite’’, ‘‘Laguerre’’, and ‘‘Jacobi’’ ensembles. These technical names reflect the fact that the p.d.f.’s for the ensembles correspond to the p.d.f.’s $\text{etr}(-A^2/2)$, $\det(A)^{a-p} \text{etr}(-A/2)$, and $\det(A)^{a_1-p} \det(I - A)^{a_2-p}$ over their respective spaces of matrices. In turn, these functions correspond to three sets of orthogonal polynomials (Hermite, Laguerre, Jacobi). Throughout this paper, we will use the term ‘‘general β -Hermite, -Laguerre, -Jacobi ensembles’’ for general β in the p.d.f.s (1), (3), (5).

Though it was believed that no other choice of β would correspond to a matrix model constructed with entries from a classical distribution, there have been studies of general β -Hermite ensembles as theoretical eigenvalue distributions. They turn out to have important applications in lattice gas theory (see [10], [2]).

The general β ensembles appear to be connected to a broad spectrum of mathematics and physics, among which we list lattice gas theory, quantum mechanics, and Selberg-type integrals. Also, the β ensembles are connected to the theory of Jack polynomials (with the correspondence $\alpha = \frac{2}{\beta}$ where α is the Jack parameter), which are currently objects of intensive research (see [27], [17], [23]).

1.3 Our Results

Dyson's original threefold way is a byproduct of the invariance assumptions as in the "Invariance" column of Figure 1. By necessity, any invariant distribution is generically dense. Further, the invariance approach forces the consideration of the complex and quaternion division algebras.

In this paper, we drop the invariance requirement. What we gain are "sparse" models (with only $O(n)$ non-zero parameters) over the reals numbers *only*. As an additional bonus, we go beyond the quantizations of the classical cases $\beta = 1, 2, 4$ and obtain continuous exponents (see Section IV for further discussion of this point).

We provide real tridiagonal random matrix models for all β -Gaussian (or Hermite) and β -Wishart (or Laguerre) ensembles, and we discuss the possibility of constructing a real matrix model for the β -MANOVA (or Jacobi) ensembles.

We obtain our results by extrapolating the classical cases, thereby providing concrete models for what have previously been considered purely theoretical distributions.

In Section II we establish results for symmetric tridiagonal matrices, and we use them to construct tridiagonal models for the β -Hermite ensembles. Along the way, we obtain a short proof based on Random Matrix Theory for the Jacobian of the transformation $T \rightarrow (q, \lambda)$, where T is a symmetric tridiagonal matrix, λ is its set of eigenvalues, and q is the first row of its eigenvector matrix. In Section III we construct tridiagonal models for the β -Laguerre ensembles, by building on the same set of ideas that we use in Section II. In Section IV we present some immediate applications of the new classes of ensembles and we discuss the β -Jacobi ensembles and other interesting open problems.

We display our random matrix constructions in Table 1.

<p>Hermite Matrix $n \in \mathbb{N}$</p>	$H_\beta \sim \frac{1}{\sqrt{2}} \begin{pmatrix} N(0, 2) & \chi_{(n-1)\beta} & & & & & \\ \chi_{(n-1)\beta} & N(0, 2) & \chi_{(n-2)\beta} & & & & \\ & & \ddots & \ddots & \ddots & & \\ & & & \chi_{2\beta} & N(0, 2) & \chi_\beta & \\ & & & & \chi_\beta & N(0, 2) & \\ & & & & & & \end{pmatrix}$
<p>Laguerre Matrix $m \in \mathbb{N}$ $a \in \mathbb{R}$ $a > \frac{\beta}{2}(m-1)$</p>	$L_\beta = B_\beta B_\beta^T, \text{ where}$ $B_\beta \sim \begin{pmatrix} \chi_{2a} & & & & & & \\ \chi_{\beta(m-1)} & \chi_{2a-\beta} & & & & & \\ & & \ddots & \ddots & & & \\ & & & \ddots & \ddots & & \\ & & & & \chi_\beta & \chi_{2a-\beta(m-1)} & \end{pmatrix}$

Table 1: Random Matrix Constructions

2 The β -Hermite (Gaussian) Ensembles

2.1 Motivation: Tridiagonalizing the GOE, GUE, and GSE

The joint distribution $f_\beta(\lambda)$ of the eigenvalues for the GOE, GUE, and GSE is

$$f_\beta(\Lambda) = c_H^\beta |\Delta(\lambda)|^\beta e^{-\frac{1}{2} \sum_i \lambda_i^2}, \quad (7)$$

where $\beta = 1, 2, 4$ [18]. Here the Vandermonde determinant notation $\Delta(\lambda)$ stands for $\prod_{i \neq j} (\lambda_i - \lambda_j)$, and c_H^β is given by (2).

We will prove in Section II.B that the tridiagonal β -Hermite random matrix displayed in Table 1 has the joint eigenvalue p.d.f. given by general β in (7). For motivation, we will begin with a quick “back-door” proof for $\beta = 1$ by tridiagonalizing the GOE; then we will extend the result to the GUE and GSE.

To illustrate the proof and help the reader follow it more easily, we have included the diagram of Figure 2.

Theorem 2.1. *If A is an $n \times n$ matrix from the GOE, then reduction of A to tridiagonal form shows that the matrix T from the 1-Hermite ensemble has joint eigenvalue p.d.f. given by (7) with $\beta = 1$.*

Proof. We write $A = \begin{pmatrix} a_n & x^T \\ x & B \end{pmatrix}$. Here a_n is a standard Gaussian, x is a vector of $(n-1)$ i.i.d. Gaussians of mean 0 and variance 1/2, and B is an $(n-1) \times (n-1)$ matrix from the GOE; a_n , x and B are all independent from each other.

Let H be any $(n-1) \times (n-1)$ orthogonal matrix (depending only on x) such that

$$Hx = [||x||_2 \ 0 \ \dots \ 0]^T \equiv ||x||_2 \ e_1,$$

where $e_1 = [1, 0, \dots, 0]^T$. Then clearly

$$\begin{pmatrix} 1 & 0 \\ 0 & H \end{pmatrix} \begin{pmatrix} a_n & x^T \\ x & B \end{pmatrix} \begin{pmatrix} 1 & 0 \\ 0 & H^T \end{pmatrix} = \begin{pmatrix} a_n & ||x||_2 e_1^T \\ ||x||_2 e_1 & HBH^T \end{pmatrix}.$$

Since A is from the GOE and H depends only on x , we can readily identify the distributions of a_n , $||x||_2$ and HBH^T (these three quantities are clearly independent). The entry a_n is unchanged and thus a standard normal with variance 1. Being the length of a multivariate Gaussian of mean 0 and entry variance 1/2, $||x||_2$ has the distribution $\frac{1}{\sqrt{2}} \chi_{n-1}$. It is worth mentioning that the p.d.f. of $||x||_2$ is given by $\frac{2}{\Gamma(\frac{n-1}{2})} y^{n-2} e^{-y^2}$. Finally, by the orthogonal invariance of the GOE, HBH^T is an $(n-1) \times (n-1)$ matrix from the GOE.

Proceeding by induction completes the tridiagonal construction.

Because the only operations we perform on A are orthogonal similarity transformations, which do not affect the eigenvalues, the conclusion of the theorem follows. \square

We recall that matrices from the GOE have the following properties:

Property 1. the joint eigenvalue density is $c_H^1 |\Delta(\lambda)| e^{-\frac{1}{2} \sum_i \lambda_i^2}$ [18];

Property 2. the first row of the eigenvector matrix is distributed uniformly on the sphere, and it is independent of the eigenvalues.

The second property is an immediate consequence of the fact that the eigenvector matrix of a GOE matrix is independent from the eigenvalues [18, (3.1.3) and (3.1.16), pages 55-58], and has the Haar (uniform) distribution because of the orthogonal invariance.

The following corollary is easily established.

Corollary 2.2. *If T is a matrix from the 1-Hermite ensemble, with eigendecomposition $T = Q\Lambda Q^T$, then the first row q of the eigenvector matrix Q is independent of Λ , and is distributed uniformly on the sphere.*

Proof. If $A = Q_1\Lambda Q_1^T$ and $T = HAH^T$, then $Q = HQ_1$. Since each one of the reflectors which form H has first row e_1 , multiplication by H does not affect the first row of Q_1 . The conclusion follows. \square

Reduction to tridiagonal form is a familiar algorithm which solves the symmetric eigenvalue problem. The special “reflector” matrix H used in practice for a vector $x = [x_1, \dots, x_{n-1}]^T$ is

$$H = I - 2\frac{uu^T}{u^T u},$$

where $u = x \pm x_1 e_1$. This special matrix H is known as the “Householder reflector” (see [11, page 209]).

The tridiagonal reduction algorithm can be applied to any real symmetric, complex hermitian, or quaternion self-dual matrix; the resulting matrix is always a real, symmetric tridiagonal. Using the algorithm similarly on a GUE or GSE matrix one gets the following

Corollary 2.3. *When $\beta = 2, 4$, reduction to tridiagonal form of matrices from the GUE, respectively GSE, shows that the tridiagonal 2-Hermite, respectively 4-Hermite, random matrix has the distribution given by (7). Note that β “counts” the number of independent Gaussians in each entry of the matrix.*

Remark 2.4. *The observation that Numerical Linear Algebra algorithms may be performed statistically is not new; it may be found in the literature (see Trotter [31], Silverstein [26], Edelman [8]).*

2.2 Tridiagonal Matrix Lemmas

In this section we prove lemmas that will be used in our constructions in Sections II.C and III.B.

Given a tridiagonal matrix T defined by the diagonal $a = (a_n, \dots, a_1)$ and sub-diagonal $b = (b_{n-1}, \dots, b_1)$, with all b_i positive, let $T = Q\Lambda Q^T$ be the eigendecomposition of T as in Theorem 2.12. Let q be the first row of Q and $\lambda = \text{diag}(\Lambda)$.

Lemma 2.5. *Under the assumptions above, starting from q and λ , one can uniquely reconstruct Q and T .*

Proof. This is a special case of the more general Theorem 7.2.1 in Parlett [24]. \square

Remark 2.6. *It follows that, except for sets of measure 0, the map $T \rightarrow (q, \lambda)$ is a bijection from the set of tridiagonal matrices of size n with positive sub-diagonal, to the set of pairs (q, λ) , with q a unit norm n -dimensional vector of positive real entries, and λ a strictly increasing ordered sequence of n real numbers. Let the bijection’s Jacobian be denoted by J ($J = \left\{ \frac{\partial(a,b)}{\partial(q,\lambda)} \right\}$).*

Our next lemma establishes a formula for the Vandermonde determinant of the eigenvalues of a tridiagonal matrix.

Lemma 2.7. *The Vandermonde determinant for the ordered eigenvalues of a symmetric tridiagonal matrix with positive sub-diagonal $b = (b_{n-1}, \dots, b_1)$ is given by*

$$\Delta(\lambda) = \prod_{i < j} (\lambda_i - \lambda_j) = \frac{\prod_{i=1}^{n-1} b_i^i}{\prod_{i=1}^n q_i},$$

where (q_1, \dots, q_n) is the first row of the eigenvector matrix.

Proof. Let $\lambda_i^{(k)}$, $i = 1 \dots k$, be the eigenvalues of the $k \times k$ lower right-corner submatrix of T . Then $P_k(x) = \prod_{i=1}^k (x - \lambda_i^{(k)})$ is the associated characteristic polynomial of that submatrix.

For $k = 1, \dots, n$ we have the three-term recurrence

$$P_k(x) = (x - a_k)P_{k-1}(x) - b_{k-1}^2 P_{k-2}(x), \quad (8)$$

and the two-term relation

$$\prod_{\substack{1 \leq i \leq k \\ 1 \leq j \leq k-1}} |\lambda_i^{(k)} - \lambda_j^{(k-1)}| = \prod_{i=1}^k |P_{k-1}(\lambda_i^{(k)})| = \prod_{j=1}^{k-1} |P_k(\lambda_j^{(k-1)})|. \quad (9)$$

From (8) we get

$$\left| \prod_{i=1}^{k-1} P_k(\lambda_i^{(k-1)}) \right| = b_{k-1}^{2(k-1)} \left| \prod_{i=1}^{k-1} P_{k-2}(\lambda_i^{(k-1)}) \right|. \quad (10)$$

By repeatedly applying (8) and (2.9) we obtain

$$\prod_{i=1}^{n-1} |P_n(\lambda_i^{(n-1)})| = b_{n-1}^{2(n-1)} \prod_{i=1}^{n-2} |P_{n-1}(\lambda_i^{(n-2)})| \quad (11)$$

$$= b_{n-1}^{2(n-1)} b_{n-2}^{2(n-2)} \left| \prod_{i=1}^{n-2} P_{n-3}(\lambda_i^{(n-2)}) \right| \quad (12)$$

$$= \dots \quad (13)$$

$$= \prod_{i=1}^{n-1} b_i^{2i}. \quad (14)$$

Finally, we use the following formula due to C.C. Paige, found in [24] as the more general Theorem 7.9.2:

$$q_i^2 = \left| \frac{P_{n-1}(\lambda_i)}{P'_n(\lambda_i)} \right| = \left| \frac{P_{n-1}(\lambda_i^{(n)})}{P'_n(\lambda_i^{(n)})} \right|. \quad (15)$$

It follows that

$$\prod_{i=1}^n q_i^2 = \frac{\prod_{i=1}^n |P_{n-1}(\lambda_i^{(n)})|}{\Delta(\lambda)^2} = \frac{\prod_{i=1}^{n-1} b_i^{2i}}{\Delta(\lambda)^2}, \quad (16)$$

which proves the result. \square

Remark 2.8. *The Vandermonde determinant formula of Lemma 2.7 can also be obtained from the Heine formula, as presented in Deift [5], page 44.*

The next lemma computes the Jacobian J by relating the tridiagonal and diagonal forms of a GOE matrix, as in Figure 2.

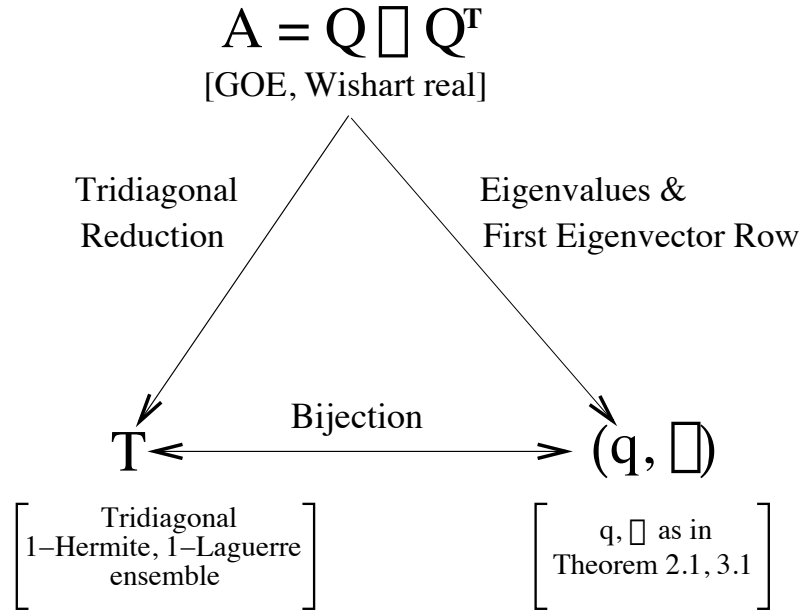


Figure 2: A dense symmetric matrix A can be tridiagonalized (left side) or diagonalized (right side). In brackets, we provide the distributions starting with that of A (GOE or Wishart real).

Lemma 2.9. *The Jacobian J can be written as*

$$J = \frac{\prod_{i=1}^{n-1} b_i}{\prod_{i=1}^n q_i}.$$

Proof. To obtain the Jacobian, we will study the transformation from GOE to 1-Hermite ensemble (see Figure 2). Note that J does *not* depend on β ; hence computing the Jacobian for this case is sufficient.

Let T be a 1-Hermite matrix. We know from Section II.A that the eigenvalues of T are distributed as the eigenvalues of a symmetric GOE matrix A , from which T can be obtained

via tridiagonal reduction ($T = HAH^T$ for some orthogonal H , which is the product of the consecutive reflections described in Section II.A).

The joint element distribution for the matrix T is

$$\mu(a, b) = c_{a,b} e^{-\frac{1}{2} \sum_{i=1}^n a_i^2} \prod_{i=1}^n b_i^{i-1} e^{-\sum_{i=1}^n b_i^2}, \quad \text{where} \quad c_{a,b} = \frac{2^{n-1}}{(2\pi)^{n/2} \prod_{i=1}^{n-1} \Gamma(\frac{i}{2})}.$$

Let

$$da = \wedge_{i=1}^n da_i, \quad db = \wedge_{i=1}^{n-1} db_i, \quad d\lambda = \wedge_{i=1}^n \lambda_i,$$

and dq be the surface element of the n -dimensional sphere. Let $\mu(a(q, \lambda), b(q, \lambda))$ be the expression for $\mu(a, b)$ in the new variables q, λ . We have that

$$\mu(a, b) da db = J \mu(a(q, \lambda), b(q, \lambda)) dq d\lambda \equiv \nu(q, \lambda) dq d\lambda. \quad (17)$$

We combine Properties 1 and 2 of Section II.A to get the joint p.d.f. $\nu(q, \lambda)$ of the eigenvalues and first eigenvector row of a GOE matrix, and rewrite it as

$$\nu(q, \lambda) dq d\lambda = n! c_H^1 \frac{2^{n-1} \Gamma(\frac{n}{2})}{\pi^{n/2}} \Delta(\lambda) e^{-\frac{1}{2} \sum_i \lambda_i^2} dq d\lambda.$$

We have introduced the $n!$ and removed the absolute value from the Vandermonde, because the eigenvalues are ordered. We have also included the distribution of q (as mentioned in Property 2, it is uniform, but only on the all-positive 2^{-n} th of the sphere because of the condition $q_i \geq 0$).

Since orthogonal transformations do not change the Frobenius norm $\|A\|_F = \sum_{i,j=1}^n a_{ij}^2$ of a matrix A , from (17), it follows that

$$J = \frac{\nu(q, \lambda)}{\mu(a, b)} = \frac{n! c_H^1 \frac{2^{n-1} \Gamma(\frac{n}{2})}{\pi^{n/2}} \Delta(\lambda)}{c_{a,b} \prod_{i=1}^n b_i^{i-1}}.$$

All constants cancel, and by Lemma 2.7 we obtain

$$J = \frac{\prod_{i=1}^{n-1} b_i}{\prod_{i=1}^n q_i}.$$

Note that we have not expressed $\mu(a, b)$ in terms of q and λ in the above, and have thus obtained the expression for the Jacobian neither in the variables q and λ , nor a and b , solely; but rather in a mixture of the two sets of variables. The reason for this is that of simplicity. \square

Remark 2.10. *Our derivation of the Jacobian is a true Random Matrix derivation. Alternate derivations of the Jacobian can be obtained either via symplectic maps or through direct calculation. We thank Percy Deift and Peter Forrester, respectively, for having shown them to us.*

Theorem 2.12. Let $H_\beta = Q\Lambda Q^T$ be the eigendecomposition of H_β ; fix the signs of the first row of Q to be non-negative and order the eigenvalues in increasing order on the diagonal of $\lambda = \text{diag}(\Lambda)$. Then λ and q , the first row of Q , are independent. Furthermore, the joint density of the eigenvalues is

$$f_\beta(\lambda) = c_H^\beta \prod_{i < j} |\lambda_i - \lambda_j|^\beta e^{-\frac{1}{2} \sum_{i=1}^n \lambda_i^2} = c_H^\beta |\Delta(\lambda)|^\beta e^{-\frac{1}{2} \sum_{i=1}^n \lambda_i^2},$$

and $q = (q_1, \dots, q_n)$ is distributed as $(\chi_\beta, \dots, \chi_\beta)$, normalized to unit length.

Proof of Theorem 2.12.

Just as before, we denote by $a = (a_n, \dots, a_1)$ the diagonal of H_β , and by $b = (b_{n-1}, \dots, b_1)$ the sub-diagonal. The differentials $da, db, dq, d\lambda$ are the same as in Lemma 2.9.

For general β , we have that

$$(dH_\beta) \equiv \mu(a, b) da db = c_{a,b} \prod_{k=1}^{n-1} b_k^{k\beta-1} e^{-\frac{1}{2} \|T_1\|_F} da db = c_{a,b} J \prod_{k=1}^{n-1} b_k^{k\beta-1} e^{-\frac{1}{2} \|T_1\|_F} dq d\lambda,$$

where

$$c_{a,b} = \frac{2^{n-1}}{(2\pi)^{n/2} \prod_{k=1}^{n-1} \Gamma\left(\frac{\beta}{2}k\right)}.$$

With the help of Lemmas 2.7 and 2.9 this identity becomes

$$(dH_\beta) = c_{a,b} \frac{\prod_{k=1}^{n-1} b_k}{\prod_{k=1}^n q_k} \prod_{k=1}^{n-1} b_k^{k\beta-1} e^{-\frac{1}{2} \|T_1\|_F} dq d\lambda \quad (21)$$

$$= c_{a,b} \frac{\prod_{k=1}^{n-1} b_k^{k\beta}}{\prod_{i=1}^n q_i^\beta} \prod_{i=1}^n q_i^{\beta-1} e^{-\frac{1}{2} \sum_i \lambda_i^2} dq d\lambda. \quad (22)$$

Thus

$$(dH_\beta) = \left(c_q^\beta \prod_{i=1}^n q_i^{\beta-1} dq \right) \left(n! c_H^\beta |\Delta(\lambda)|^\beta e^{-\frac{1}{2} \sum_i \lambda_i^2} d\lambda \right).$$

Since the joint density function of q and λ separates, q and λ are independent. Moreover, once we drop the ordering imposed on the eigenvalues, it follows that the joint eigenvalue density of H_β is $c_H^\beta |\Delta(\lambda)|^\beta e^{-\frac{1}{2} \sum_i \lambda_i^2}$, and q is distributed as $(\chi_\beta, \dots, \chi_\beta)$, normalized to unit length. From (22), it also follows that

$$c_q^\beta = \frac{2^{n-1} \Gamma\left(\frac{\beta}{2}n\right)}{\left[\Gamma\left(\frac{\beta}{2}\right)\right]^n}. \quad (23)$$

□

3 The β -Laguerre (Wishart) Ensembles

3.1 Motivation: Tridiagonalizing the Wishart ensembles

The preceding section gives tridiagonal random matrix models for all β -Hermite ensembles. In the following we define the β -Laguerre ensembles, and give tridiagonal random matrix models for them.

The Wishart ensembles have joint eigenvalue density

$$f_{\beta}(\lambda) = c_L^{\beta,a} |\Delta(\lambda)|^{\beta} \prod_{i=1}^m \lambda_i^{a-p} e^{-\sum_{i=1}^m \lambda_i/2}, \quad (24)$$

again with $a = \frac{\beta}{2}n$, $p = 1 + \frac{\beta}{2}(m-1)$, and with $\beta = 1$ for real, respectively $\beta = 2$ for complex. Here $c_L^{\beta,a}$ as the same as in (4).

From now on p will always denote the quantity $1 + \frac{\beta}{2}(m-1)$, following the notation of Muirhead for $\beta = 1$ [21] (chapter 7) and Forrester [10] (Forrester uses $1 + \frac{1}{\alpha}(m-1)$, where $\alpha = 2/\beta$ is the Jack parameter). Its presence is implicit in the p.d.f. of all β -Laguerre ensembles; hence we will identify the ensembles by β and by a (we call the latter the ‘‘Laguerre’’ parameter, generalizing from the univariate case $\beta = 1$, $m = 1$).

As in Section II.A, we will provide the most basic case for our construction: the case $\beta = 1$ and Wishart real exponent $\frac{n-m-1}{2}$ (also referred as the case $\beta = 1$ and Laguerre parameter $a = \frac{n}{2}$).

Theorem 3.1. *Let G be an $m \times n$ matrix of i.i.d. standard Gaussians; then $W = GG^T$ is a Wishart real matrix. By reducing G to bidiagonal form B one obtains that the matrix $T = BB^T$ from the 1-Laguerre ensemble of Laguerre parameter $a = \frac{n}{2}$ (defined as in Table 1) has the joint eigenvalue p.d.f. given by (24).*

Proof. We write

$$G = \begin{pmatrix} x^T \\ G_1 \end{pmatrix},$$

with x^T a row multivariate standard Gaussian of length n and G_1 a $(m-1) \times n$ matrix of i.i.d. standard Gaussians. Let R be a right reflector corresponding to the vector x^T ($R^T x = \|x\|_2 e_1^T$) which is independent of G_1 . Hence $G_1 R$ is a matrix of i.i.d. standard Gaussians.

Write $G_1 R = [y, G_2]$, where y is a column multivariate standard Gaussian of length $m-1$ and G_2 is a $(m-1) \times (n-1)$ matrix of i.i.d. standard Gaussians. Let L be a left reflector corresponding to y ($Ly = \|y\|_2 e_1$) which is independent of G_2 . Then we have that

$$\begin{pmatrix} 1 & 0 \\ 0 & L \end{pmatrix} G R = \begin{pmatrix} \|x\|_2 & 0 \\ \|y\|_2 e_1 & LG_2 \end{pmatrix}.$$

As we have seen before, $\|x\|_2$ is distributed like χ_{n-1} , $\|y\|_2$ is distributed like χ_{m-1} , and LG_2 is a matrix of i.i.d. standard Gaussians (since L and G_2 are independent).

We proceed inductively to finish the bidiagonal construction of B .

Because the operations we have performed on G are orthogonal left and right multiplications, which do not affect the singular values, it follows that the singular values of G and B are the same. Since the squares of the singular values of G , respectively B , are the eigenvalues of W , respectively T , the conclusion of the theorem follows. \square

Remark 3.2. *The bidiagonalization process presented above is part of a familiar Numerical Linear Algebra algorithm for computing the singular values of a matrix.*

Corollary 3.3. *The same process of bidiagonalization performed on \tilde{G} , a matrix of i.i.d. standard complex (standard quaternion) Gaussians, shows that the matrix $\tilde{W} = \tilde{G}\tilde{G}^T$ and the matrix T from the 2-Laguerre (4-Laguerre) ensemble of parameter $a = n$ ($a = 2n$) has the joint eigenvalue p.d.f. given by (24). In all three cases (real, complex, quaternion) we say that T represents the tridiagonalization of the Wishart (real, complex, quaternion) ensemble.*

In the next section we prove the general form of the theorem.

3.2 The Eigendistribution of β -Laguerre Ensemble

Let

$$B_\beta \sim \begin{pmatrix} \chi_{2a} & & & & \\ \chi_{\beta(m-1)} & \chi_{2a-\beta} & & & \\ & \ddots & \ddots & & \\ & & & \chi_\beta & \chi_{2a-\beta(m-1)} \end{pmatrix},$$

by this meaning that all of the $2m - 1$ diagonal and subdiagonal elements are mutually independent with the corresponding χ distribution.

Let $L_\beta = B_\beta B_\beta^T$ be the corresponding tridiagonal matrix.

Theorem 3.4. *Let $L_\beta = Q\Lambda Q^T$ be the eigendecomposition of L_β ; fix the signs of the first row of Q to be non-negative and order the eigenvalues increasingly on the diagonal of Λ . Then Λ and the first row q of Q are independent. Furthermore, the joint density of the eigenvalues is*

$$f_\beta(\lambda) = c_L^{\beta,a} |\Delta(\lambda)|^\beta \prod_{i=1}^n \lambda_i^{a-p} e^{-\sum_{i=1}^n \lambda_i/2},$$

where $p = 1 + \frac{\beta}{2}(m - 1)$, and q is distributed as $(\chi_\beta, \dots, \chi_\beta)$ normalized to unit length.

Proof of Theorem 3.4. We will use throughout the results of Lemma 2.7, Lemma 2.9, Lemma 2.11 and Remark 2.6, which are true in the context of tridiagonal symmetric matrices with positive sub-diagonal entries. By definition, L_β is such a matrix.

We will again use the notations of Lemma 2.9 and 2.11 for the differentials da , db , dq , $d\lambda$, dx , and dy .

We define (dB_β) to be the joint element distribution on B_β

$$(dB_\beta) \equiv \mu(x, y) dx dy = \prod_{i=0}^{m-1} x_{m-i}^{a-\beta i-1} e^{-x_i^2/2} \prod_{i=1}^{m-1} y_i^{\beta i-1} e^{-y_i^2/2} dx dy.$$

By using Lemma 24 we obtain the joint element distribution on L_β as

$$(dL_\beta) \equiv J_{B \rightarrow T}^{-1} \mu(x, y) dx dy \tag{25}$$

$$= 2^{-m} c_{x,y} x_1^{2a-\beta(m-1)-2} e^{-x_1^2/2} \prod_{i=0}^{m-2} x_{m-i}^{a-\beta i-3} e^{-x_i^2/2} \prod_{i=1}^{m-1} y_i^{\beta i-1} e^{-y_i^2/2} dx dy, \tag{26}$$

where

$$c_{x,y} = \frac{\prod_{i=1}^{m-1} \Gamma(i\frac{\beta}{2}) \prod_{i=1}^m \Gamma(a - \frac{\beta}{2}(i-1))}{2^{2m-1}}.$$

We rewrite (26) in terms of x, y, λ , and q :

$$\begin{aligned} (dL_\beta) &= 2^{-m} c_{x,y} e^{-\sum_{i=1}^m x_i^2/2} e^{-\sum_{i=1}^{m-1} y_i^2/2} \frac{\prod_{i=1}^{m-1} (x_{i+1} y_i)}{\prod_{i=1}^m q_i} x_1^{2a-\beta(m-1)-2} \times \\ &\quad \times \prod_{i=0}^{m-2} x_{m-i}^{2a-\beta(m-i)-3} \prod_{i=1}^{m-1} y_i^{\beta i-1} dq d\lambda \\ &= 2^{-m} c_{x,y} e^{-\sum_{i=1}^m x_i^2/2} e^{-\sum_{i=1}^{m-1} y_i^2/2} \frac{\prod_{i=0}^{m-1} x_{m-i}^{2a-\beta(m-i)-2} \prod_{i=1}^{m-1} y_i^{\beta i}}{\prod_{i=1}^m q_i} dq d\lambda. \end{aligned}$$

Since the Vandermonde with respect to b and q and the ordered eigenvalues λ can be written as

$$\Delta(\lambda) = \frac{\prod_{i=1}^{m-1} b_i}{\prod_{i=1}^m q_i},$$

it follows that

$$\Delta(\lambda) = \frac{\prod_{i=1}^{m-1} (x_{i+1} y_i)^i}{\prod_{i=1}^m q_i}.$$

This means that we can rewrite

$$\begin{aligned} (dL_\beta) &= 2^{-m} c_{x,y} e^{-\sum_{i=0}^{m-1} x_{n-i}^2/2} e^{-\sum_{i=1}^{m-1} y_i^2/2} \frac{\prod_{i=1}^{m-1} (x_{i+1} y_i)^{\beta i}}{\prod_{i=1}^m q_i^\beta} \prod_{i=1}^{m-1} q_i^{\beta-1} \prod_{i=0}^{m-1} x_{m-i}^{2a-\beta(m-1)-2} dq d\lambda \\ &= 2^{-m} c_{x,y} e^{-\sum_{i=0}^{m-1} x_{n-i}^2/2} e^{-\sum_{i=1}^{m-1} y_i^2/2} \Delta(\lambda)^\beta \prod_{i=1}^{m-1} q_i^{\beta-1} \left(\prod_{i=0}^{m-1} x_{m-i} \right)^{2a-\beta(m-1)-2} dq d\lambda. \end{aligned}$$

The trace and the determinant are invariant under orthogonal similarity transformations, so $\text{tr}(L_\beta) = \text{tr}(\Lambda)$, and $\det(L_\beta) = \det(\Lambda)$. This is equivalent to

$$\sum_{i=0}^{m-1} x_{m-i}^2 + \sum_{i=1}^{m-1} y_i^2 = \sum_{i=1}^m \lambda_i,$$

$$\prod_{i=0}^{m-1} x_{m-i}^2 = \prod_{i=1}^m \lambda_i.$$

Using this, and substituting p for $1 + \frac{\beta}{2}(m-1)$, we obtain that

$$(dL_\beta) = \left(c_q^\beta \prod_{i=1}^{m-1} q_i^{\beta-1} dq \right) \left(m! c_L^{\beta,a} e^{-\sum_{i=1}^m \lambda_i/2} \Delta(\lambda)^\beta \prod_{i=1}^m \lambda_i^{a-p} d\lambda \right),$$

where c_q^β is the same as in (23).

From the above we see that q and λ are independent, and once we drop the ordering the joint eigenvalue density is given by the β -Laguerre ensemble of parameter a , while q is distributed like a normalized vector of χ_β 's.

This concludes the proof of Theorem 3.4. □

4 Applications and Open Problems

As we mentioned in Section 1, we believe that there should be many applications for the new tridiagonal ensembles. We illustrate here some (in Section IV.A), in the hope that researchers will find many more. Some of the applications we believe are new results (Applications 1, 3, 5, and 6), and some are simplifications of known results (Applications 2 and 4).

We discuss the open problem of constructing a matrix model for the β -Jacobi ensembles in the beginning of Section IV.B. To facilitate the finding of new results, we conclude with a few open “general β -ensemble” problems.

4.1 Applications

1. *Interpolating Laguerre exponents.* Our β -Laguerre ensembles have “continuous” Laguerre parameters a which, even in the cases $\beta = 1, 2, 4$, interpolate the Wishart parameters.

Though β -Laguerre ensembles with general (“continuous”) parameter a have been studied by many researchers ([2], [14], [21]), no non-quantized matrix realizations (i.e. explicit random matrix models) of β -Laguerre ensembles are found in the literature.

By “quantized” we mean that the exponent a is either an even integer, an integer, or a half-integer (depending on the value of β). In particular, all models corresponding to a Laguerre (or Jacobi) weight found in [33] and [12] are quantized.

Thus, our β -Laguerre random matrix constructions extend the pre-existing ones in two ways: through β and through the Laguerre parameter a .

2. *The expected characteristic polynomial.* The result below might be seen as an extension of the classical Heine theorem (see Szegő [25] and Deift [5]) which has $\beta = 2$. Note that for $\beta \neq 2$, $\Delta(\lambda)^\beta$ can no longer be written as the determinant of a Vandermonde matrix times its transpose, and the proof cannot be duplicated.

The same result is found in a slightly more general form in [8], and its Jacobi case was first derived by Aomoto [1].

Theorem 4.1. *The expected characteristic polynomial $P_n(y) = \det(yI_n - S)$ over S in the β -Hermite, respectively β -Laguerre, ensembles are proportional to*

$$H_n\left(\frac{y}{\sqrt{2\beta}}\right), \quad \text{respectively to} \quad L_n^{\frac{2a}{\beta}-n}\left(\frac{y}{2\beta}\right).$$

Here H_n and $L_n^{\frac{2a}{\beta}-n}$ are the Hermite, respectively Laguerre, polynomials, and the constant of proportionality accounts for the fact that $P_n(y)$ is monic.

Proof. Both formulas follow immediately from the 3-term recurrence for the characteristic polynomial of a tridiagonal matrix (see formula (8)) and from the independence of the variables involved in the recurrence. \square

3. *Expected values of symmetric polynomials.* Using the three-term recurrence for the characteristic polynomial of a tridiagonal matrix, we obtain Theorem 4.2.

Theorem 4.2. *Let p be any fixed (independent of β) multivariate symmetric polynomial on n variables. Then the expected value of p over the β -Hermite or β -Laguerre ensembles is a polynomial in β .*

We remark that it is difficult to see this from the eigenvalue density.

Proof. The elementary symmetric functions

$$e_i(x_1, x_2, \dots, x_n) = \sum_{1 \leq j_1 < \dots < j_i \leq n} x_{j_1} x_{j_2} \dots x_{j_i} \quad i = 0, 1, \dots, n,$$

can be used to generate any symmetric polynomial of degree n (in particular p).

The e_i evaluated at the eigenvalues of a matrix are the coefficients of its characteristic polynomial, and hence they can be written in terms of the matrix entries. Thus p can be written as a polynomial of the $n \times n$ tridiagonal matrix entries (which corresponds to the Hermite, respectively Laguerre cases).

To obtain the expected value of p over the β -Hermite or β -Laguerre ensemble, one can write p in terms of the corresponding matrix entries, use the symmetry to condense the expression, then replace the powers of the matrix entries by their expected values.

The diagonal matrix entries are either normal random variables in the Hermite case or sums of χ^2 random variables in the Laguerre case. The subdiagonal entries appear only raised at even powers in the e_i and hence in p (this is an immediate consequence of the three-term recurrence for the characteristic polynomial, (8)). Since all even moments of the involved χ distributions are polynomials in $\beta/2$, it follows that the expectation of p will be a polynomial in β . \square

As an easy consequence we have the following corollary.

Corollary 4.3. *All moments of the determinant of a β -Hermite matrix are **integer-coefficient** polynomials in $\beta/2$.*

Proof. Note that even moments of the $\chi_{\beta i}$ distribution are integer-coefficient polynomials in $\beta/2$, and that the determinant is e_n . \square

4. *A new proof for Hermite and Laguerre forms of the Selberg Integral.* Here is a quick proof for the Hermite and Laguerre forms of the Selberg Integral ([18]), using the β -Hermite, respectively, β -Laguerre ensembles.

The Hermite Selberg integral is

$$I_H(\beta, n) \equiv \int_{\mathbb{R}^n} |\Delta(\lambda)|^\beta e^{-\sum_{i=1}^n \lambda_i^2/2} d\lambda$$

We have that

$$I_H(\beta, n) = n! \left(\int_{0 \leq \lambda_1 \leq \dots \leq \lambda_n < \infty} \Delta(\lambda)^\beta e^{-\sum_{i=1}^n \lambda_i^2/2} d\lambda \right) \left(c_q^\beta \int_{S_+^{n-1}} \prod_{i=1}^n q_i^{\beta-1} dq \right),$$

where c_q^β is as in (23). We introduce the $n!$ because in the first integral we have ordered the eigenvalues; S_+^{n-1} signifies that all q_i are positive.

Note that c_q^β can easily be computed independently of the β -Hermite ensembles.

Using the formula for the Vandermonde given by Lemma 2.7, the formula for the Jacobian J given in Lemma 2.9, and the fact that the Frobenius norm of a matrix in the tridiagonal

1-Hermite ensemble is the same as the Frobenius norm of its eigenvalue matrix, one obtains

$$\begin{aligned}
I_H(\beta, n) &= n! c_q^\beta \int_{\mathbb{R}^n \times (0, \infty)^{n-1}} \frac{\prod_{i=1}^n q_i \prod_{i=1}^{n-1} b_i^{\beta i}}{\prod_{i=1}^{n-1} b_i \prod_{i=1}^n q_i^\beta} \prod_{i=1}^n q_i^{\beta-1} e^{-\sum_{i=1}^n b_i^2 - \sum_{i=1}^n a_i^2/2} da db \\
&= n! c_q^\beta (2\pi)^{n/2} \prod_{i=1}^{n-1} \int_{(0, \infty)} b_i^{\beta i-1} e^{-b_i^2} db_i \\
&= n! \frac{2^{n-1} \Gamma(\frac{\beta}{2}n)}{\left(\Gamma(\frac{\beta}{2})\right)^n} (2\pi)^{n/2} \prod_{i=1}^{n-1} \frac{\Gamma(\frac{\beta}{2}i)}{2} = \frac{1}{c_H^\beta}.
\end{aligned}$$

The same reasoning yields the Laguerre Selberg Integral formula

$$I_L^{\beta, a, n} = \frac{1}{c_L^{\beta, a}}.$$

5. *Moments of the discriminant.* The discriminant of a polynomial equation of order m is the square of the Vandermonde determinant of the m zeroes of the equation. Thus, the discriminant of the characteristic polynomial of a β -Hermite or β -Laguerre ensemble matrix is simply $D(\lambda) = \Delta(\lambda)^2$.

A simple calculation shows that the k th moment of $D(\lambda)$ is

$$\begin{aligned}
\frac{c_H^\beta}{c_H^{\beta+2k}} &= \prod_{j=1}^n \frac{(1 + \frac{\beta}{2}j)_{kj}}{(1 + \frac{\beta}{2})_k}, \quad \text{respectively,} \\
\frac{c_L^{\beta, a}}{c_L^{\beta+2k, a+k(m-1)}} &= 2^{km(m-1)} \prod_{j=1}^m \frac{(1 + \frac{\beta}{2}j)_{kj} (a - \frac{\beta}{2}(m-j))_{k(j-1)}}{(1 + \frac{\beta}{2})_k}.
\end{aligned}$$

where n and m are the matrix sizes for the Hermite, respectively, Laguerre cases, and the rising factorial $(x)_k \equiv \Gamma(x+k)/\Gamma(x)$.

Using the Selberg integral, one obtains that the moments of the discriminant for the β -Jacobi case are

$$\frac{c_J^{\beta, a_1, a_2}}{c_J^{\beta+2k, a_1+k(m-1), a_2+k(m-1)}} = \prod_{j=1}^m \frac{(1 + \frac{\beta}{2}j)_{kj} (a_1 - \frac{\beta}{2}(m-j))_{k(j-1)} (a_2 - \frac{\beta}{2}(m-j))_{k(j-1)}}{(1 + \frac{\beta}{2})_k (a_1 + a_2 - \frac{\beta}{2}(m-j))_{k(m+j-2)}}.$$

6. *Software for Application 2. Computing eigenvalue statistics for the β -ensembles.* Application 2 suggests that integrals of the form

$$E_\beta[p] \equiv c_H^\beta \int_{\mathbb{R}^n} p(\lambda) |\Delta(\lambda)|^\beta e^{-\sum_{i=1}^n \lambda_i^2/2} d\lambda$$

may be evaluated with software.

One example of this would be computing moments of the determinant over the β -Hermite ensemble. There are explicit formulas for the cases $\beta = 1, 2$ and 4 , due to Mehta [19] and to Delannay and Le Caër [6], which can be used to evaluate these moments.

In the absence of a closed-form, explicit formula, like the one for $\beta = 1$ provided in [6], the computation of these moments cannot be made polynomial; thus it is inherently slow.

For the general β case, one can compute the moments in terms of a multivariate Hermite polynomial evaluated at 0 (see [4], [2]). Using this technique, the complexity of the computation might exceed that of symbolically taking the determinant of a tridiagonal matrix, expanding the power, and replacing all powers of the entries by their expected values (which are all known). Writing a Mathematica code to implement this algorithm is an easy exercise, and such a code would allow the author to compute these moments in a reasonable amount of time, provided that the product between the power and the size of the matrix is not very large. A template for a special case when $\beta = 1$ can be found in [9, Appendix A].

4.2 Open Problems

1. *β -Jacobi (MANOVA) Ensembles.* Sections II and III of the paper provide tridiagonal matrix models for the β -Hermite and β -Laguerre ensembles. The natural question is whether such models exist for the last member of the classical triplet, Jacobi. The β -Jacobi ensembles have been intensively studied as theoretical distributions, especially in connection with Selberg-type integrals and Jack (or Jack-Selberg) polynomials (see [1], [15], [16], [3]). Finding a random matrix model that corresponds to them would be of much interest.

If the two matrix factorizations problems that are associated with the Hermite and Laguerre ensembles are the EIG and the SVD, the one associated with the Jacobi should be the QZ (the generalized symmetric eigenvalue problem). This idea is supported by the fact that the MANOVA real and complex distributions, which correspond to the Jacobi $\beta = 1, 2$ ensembles, are indeed connected to the QZ algorithm. A good reference for QZ is [11].

Though we have not studied this problem sufficiently, we believe that a concrete (perhaps sparse, perhaps tridiagonal) matrix model may be constructed for the β -Jacobi ensembles.

2. *Level densities.* The level density of an ensemble is the distribution of a random eigenvalue of that ensemble (and by the Wigner semicircular law we know that the limiting distribution as $n \rightarrow \infty$ of such an eigenvalue is semicircular). The three functions found to be the level densities of the Gaussian models depend on the univariate Hermite polynomials.

Recently, Forrester [10] has found a formula for the level densities of the β -Hermite ensembles which works for β an even integer. This formula depends on a multivariate Hermite polynomial.

Finding a unified formula for the general β case would be of interest.

3. *Level spacings.* The level spacings are the distances between the eigenvalues of an ensemble, usually normalized so that the average consecutive spacing is 1. These spacings have been well-studied in the case of the Gaussian ensembles ($\beta = 1, 2, 4$). The limiting probability density of a random spacing in these cases is known in terms of spheroidal functions (see [18]).

A surprising connection exists between the limiting probability density of a GUE random spacing and the probability density of the zeroes of the Riemann zeta function. Inspired by the theoretical work of Montgomery ([20]), Odlyzko ([22]) has shown experimentally that the two probability densities are very close; the subsequent conjecture that the two probability densities coincide has been named the Montgomery-Odlyzko law.

To the best of our knowledge, the level spacing of the general β -Hermite ensembles has not been investigated.

4. *Bulk and edge scaling limits.* Finally, a very important application would be the generalization of the bulk and edge scaling limits for the GOE, GUE and GSE obtained by Tracy and Widom (the latter are known as the Tracy-Widom distributions F_1 , F_2 and F_4).

The edge scaling limit refers to the distribution of the largest eigenvalue of a matrix in the ensemble; the bulk scaling limit refers to the distribution of an eigenvalue in the “bulk” of the spectrum. For a reference, see [29], [30] or [28]. The Tracy-Widom distributions are defined in terms of Painlevé functions, which are solutions to certain differential equations, with asymptotics given by Airy functions. For a good treatment of Painlevé equations in relationship with Gaussian (Hermite), Laguerre, and Jacobi random matrix models, see Pierre van Moerbeke’s notes [32, Section 4]. Recently, Johnstone [14] has found that the limiting distributions F_1 and F_2 apply to real (respectively complex) Wishart matrices.

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